A Little Optimization

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Logistic Regresssion

Consider the linear regression model with iid normal errors:

$$Y = X \beta + \epsilon, \ \epsilon \sim N(0, \sigma^2 I)$$

We know how to get the least squares estimate of β which is also the mle. We solve the equations:

$$X'(Y-X\beta)=0$$

We can solve these equation directly using $\hat{\beta} = (X'X)^{-1}X'Y$ or use the cholesky decomposition of X'X or the QR decomposition of X.

We have also learned how to do a Bayesian analysis of the linear model using the normal prior for β and the inverted chi-squared prior for σ and the Gibbs sampler:

 $\beta \mid \sigma, (Y, x), \sigma \mid \beta, (Y, X)$

Perhaps after linear regression, the most fundamental model in applied statistics is logistic regression.

We want to use linear methods, but now are response is binary: $y \in \{0, 1\}$. For a single (x, y) observation we have

$$P(Y = 1 | x, \beta) = F(x' \beta), \ F(\eta) = \frac{\exp(\eta)}{1 + \exp(\eta)}.$$

How do we compute the mle???

How do we do a Bayesian analysis

Newton's Method

Newton's method is a very basic method in optimization.

We will use it to compute the logit mle, and the Bayesian posterior mode.

Suppose $f : \beta \to R$, $\beta \in R^p$.

We want to minimize (or maximize) f.

We will need the first derivative:

$$f'(\beta) = \nabla f(\beta) = \left[\frac{\partial f(\beta)}{\partial \beta_1}, \frac{\partial f(\beta)}{\partial \beta_2}, \dots, \frac{\partial f(\beta)}{\partial \beta_p}\right]$$

and the second derivative:

$$f''(\beta) = \left[\frac{\partial^2 f(\beta)}{\partial \beta_i \partial \beta_j}\right]$$

- the first derivative is called the gradient vector. My convention is that it is a 1 × p row vector.
- the second derivative is a p × p symmetric matrix. It is called the Hessian.

Newton's Method

Newton's method is iterative.

Let β_i the value at iteration *i*.

- approximate f at β_i by a quadratic using Taylors's theorem.
- optimize the quadratic: the solution is β_{i+1} .
- repeat until converged.

Taylor approximation:

$$f(eta) pprox ilde{f}(eta) = f(eta_i) + f'(eta_i)(eta - eta_i) + rac{1}{2}(eta - eta_i)'f''(eta_i)(eta - eta_i)$$

Now to optimize the quadratic, we compute its gradient and set it equal to 0.

$$abla ilde{f}(eta) = f'(eta_i) + (eta - eta_i)' f''(eta_i)$$

We can solve $\nabla \tilde{f}(\beta) = 0$ with $0 = f'(\beta_i) + (\beta - \beta_i)'f''(\beta_i)$ $-f'(\beta_i)[f''(\beta_i)]^{-1} = \beta' - \beta'_i$ $\beta' = \beta'_i - f'(\beta_i)[f''(\beta_i)]^{-1}$

$$\beta_{i+1} = \beta_i - [f''(\beta_i)]^{-1} [f'(\beta_i)]'$$

Logit Log-Likelihood Derivatives: The Logit MLE

We will compute the first and second derivatives of the logit log likelihood.

First, we differentiate $F(\eta) = \frac{\exp(\eta)}{1 + \exp(\eta)}$:

$$F'(\eta) = rac{(1+e^{\eta})e^{\eta} - e^{\eta}e^{\eta}}{(1+e^{\eta})^2} = F(\eta)(1-F(\eta))$$

The Likelihood

$$L(\beta) = \prod_{i=1}^{n} F(x'_{i}\beta)^{y_{i}} (1 - F(x'_{i}\beta))^{(1-y_{i})}$$

Let $F_i = F(x'_i\beta)$.

$$\log L(\beta) = \sum y_i \, \log(F_i) + (1 - y_i) \log(1 - F_i)$$

$$\log \mathcal{L}'(\beta) = \sum y_i x'_i \frac{F_i(1-F_i)}{F_i} + x'_i(1-y_i) \left[-\frac{F_i(1-F_i)}{(1-F_i)}\right]$$

= $\sum [y_i x'_i(1-F_i) - x'_i(1-y_i)F_i]$
= $\sum x'_i(y_i - F_i)$
= $(y - F)'X$

$$\log L''(\beta) = -\sum_{i} x_i x_i' F_i (1 - F_i)$$
$$= -X' DX$$

where,

$$D = diag(F_i(1 - F_i))$$

So, to compute the logit mle:

$$\beta_{i+1} = \beta_i - [-X'DX]^{-1}X'(y - F) = \beta_i + [X'DX]^{-1}X'(y - F)$$

Iteratively Reweighted Least Squares

Recall weighted least squares

$$Y = X\beta + \epsilon, \ \epsilon \sim N(0, \Sigma)$$

then,

$$\hat{\beta} = (X'\Sigma^{-1}X)^{-1}X'\Sigma^{-1}y$$

It may be helpful to rewrite the Newton iteration as a series of weighted regressions:

Let $\Sigma^{-1} = D$ and

$$Z = X\beta_i + D^{-1}(y - F)$$

then,

$$(X'\Sigma^{-1}X)^{-1}X'\Sigma^{-1}Z = (X'\Sigma^{-1}X)^{-1}X'\Sigma^{-1}(X\beta_i + D^{-1}(y - F))$$

= $\beta_i + [X'DX]^{-1}X'(y - F)$

Hence doing an iteratively (re)weighted least squares problem (IRLS) gets you the mle.

Optimization and Convexity/Concavity

Recall that a function is convex if

$$f(\alpha x_1 + (1 - \alpha)x_2) \le \alpha f(x_1) + (1 - \alpha)f(x_2), \ \alpha \in [0, 1].$$

and concave if it goes the the other way,

$$f(\alpha x_1 + (1 - \alpha)x_2) \ge \alpha f(x_1) + (1 - \alpha)f(x_2), \ \alpha \in [0, 1].$$

If a function is convex then it has a unique global minimum and any local minimum is the local minimum.

Same for concave and maximum.

Key:

If the Hessian is positive definite everywhere, then the function is convex.

Same for negative definite and concave.

$$a'[-X'DX]a = -v'Dv = -\sum v_i^2 F_i(1-F_i) \le 0.$$

Hence the logit logLikelihood is concave, hence Newton will converge to a global max.

Bayesian Posterior Mode

We can use a very similar approach to compute the Bayesian posterior mode given a mulitvariate normal prior for β .

Let

$$p(\beta) \sim N(\overline{\beta}, A^{-1}).$$

Then,

 $p(\beta | X, Y) \propto L(\beta) p(\beta)$

and

$$\log p(\beta \,|\, X, Y) = \log L(\beta) + \log(p(\beta))$$

$$log(p(\beta)) = C - rac{1}{2}(eta - areta)'A(eta - areta) \equiv C + g(eta)$$
 $g'(eta) = -(eta - areta)A, \ g''(eta) = -A.$

So the Newton iterations become:

$$\beta_{i+1} = \beta_i + [X'DX + A]^{-1}[X'(y - F) - A(\beta_i - \bar{\beta})]$$

where D and F also depend on β_i .

Note:

Maximizing the log posteior is equivalent to minimizing

$$-\log L(\beta) + \frac{1}{2}(\beta - \overline{\beta})'A(\beta - \overline{\beta})$$

If we let $A = \lambda I$, $\bar{\beta} = 0$, and recall that $-\log L(\beta)$ is the deviance which is also called the cross-entropy loss then we minimize

$$Loss(y, \beta) + \lambda ||\beta||^2$$

Thus the Bayesian posterior mode can be viewed as an L2 regularized estimate of the coefficient vector.